Terms of Trade and Total Factor Productivity: Empirical evidence from Latin American emerging markets

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Introduction

- In the last decade high terms of trade have boosted economic activity in many emerging market economies
- In the region Mexico, Chile, Peru are among the economies that have benefited from this favorable external environment
- Typically, this economies have experimented a boost in investment in the tradable sector, which have increased potential output, and also TFP.
- How much of the increase in TPF can be linked to high terms of trade?
- In this paper we present a two step procedure to answer this question:
 - ullet First we use a DSGE model o to estimated TFP: Signal extraction stage
 - $\{ToT_t, TFP_t\}_{t=0}^{\infty} \to TFP$ decomposition between domestic and external factors: SVAR stage.
- We use quarterly data for Chile, Mexico and Peru.

Features of the Model

- Our framework follows a standard RBC model for a small open economy similar to Chang and Fernández (2013) and García-Cicco et.al. The model main features are: (2010).
- Imperfect access to international capital markets, which generates an endogenous risk premium, linked to the foreign net debt position.
- Variable capacity utilization. Firms can expand output not only by hiring more workers but also by using capital more intensively, at the cost of accelerating depreciation.
- Permanent productivity shock. Variables in the model are not stationary.

Main equations

Production function

$$Y_t = a_t (U_t K_t)^{\alpha} (A_t N_t)^{1-\alpha}$$

Capital Accumulation

$$K_{t+1} = I_t + (1 - \delta U_t^{\phi}) K_t - \frac{\psi_K}{2} \left(\frac{K_{t+1}}{K_t} - X \right)^2 K_t,$$

• Interest rate, country-specific spread, and the risk premium

$$\begin{array}{rcl} R_t & = & R_t^* S_t + \psi_D \left(e^{\widetilde{D}_{t+1} - D} - 1 \right), \\ ln(S_t / S) & = & \eta \mathbb{E}_t (ln X_{t+1} + a_{t+1}), \\ ln(R_t^* / R^*) & = & \rho_{r^*} ln(R_{t-1}^* / R^*) + \varepsilon_t^{r^*}, \varepsilon_t^{r^*} \sim N(0, \sigma^{r^*}), \end{array}$$

Trade Balance

$$Y_t - C_t - I_t = D_t - \frac{D_{t+1}}{R_t}$$
,

Main equations

• Euler equation and labor market equilibrium condition

$$\begin{array}{rcl} 1 & = & \beta E_t \left(R_t \frac{U_{c,t+1}}{U_{c,t}} \right), \\ \tau v A_{t-1} N_t^{\upsilon - 1} & = & (1-\alpha) \frac{Y_t}{N_t}, \end{array}$$

where $U_{c,t} = [C_t/A_{t-1} - \tau N_t^{v}]^{-\sigma}$,

• Investment decision

$$1 + \psi_k \mathbb{E}_t \left(\frac{K_{t+1}}{K_t} - X \right) = \beta \mathbb{E}_t \left\{ \frac{U_{c,t+1}}{U_{c,t}} \left[\alpha \frac{Y_{t+1}}{K_t} + 1 - \delta U_{t+1}^{\phi} + \frac{\psi_k}{2} \left(\left(\frac{K_{t+2}}{K_{t+1}} \right)^2 - X^2 \right) \right] \right\},$$

• Capital utilization decision

$$\alpha \frac{Y_t}{U_t} = \varphi \delta U_t^{\varphi - 1} K_t.$$

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Main equations

The TFP, $a_t(A_t)^{1-\alpha}$, has to components:

• The permanent productivity shock A_t , that is assumed to follow a random walk in logs:

$$\begin{split} \ln & A_t &= \ln A_{t-1} + \ln X_t, \\ \ln & X_t &= (1 - \rho_x) \ln(X) + \rho_x \ln X_{t-1} + \varepsilon_t^x, \varepsilon_t^x \sim N(0, \sigma^x). \end{split}$$

• The transitory productivity shock, a_t , that follows a stationary autoregressive stochastic process in logs:

$$lna_t = \rho_a lna_{t-1} + \varepsilon_t^a, \varepsilon_t^a \sim N(0, \sigma^a).$$

 The model is calibrated to reproduce main business cycle features of the three economies analysed.

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Moments

Table: Second moments

Statistics	Statistics : Mexico							
		Data				Mo	del	
$\Delta ln(Y)$	$\Delta ln(C)$	$\Delta ln(I)$	$\Delta TB/Y$	$\Delta ln(ToT)$	$\Delta ln(Y)$	$\Delta ln(C)$	$\Delta ln(I)$	$\Delta TB/Y$
Standard	Standard deviation (in %)							
1.645	1.981	5.104	1.269	4.874	1.884	1.920	5.791	1.335
Relative .	$S.D w.r.t \Delta l$	n(Y)			•			
1.000	1.205	3.103	0.772	2.963	1.000	1.019	3.073	0.709
Correlati	on with Δlr	$\iota(Y)$			•			
1.000	0.827	0.742	-0.374	0.255	1.000	0.969	0.905	-0.672
Correlati	on with ΔT	B/Y			•			
-0.374	-0.431	-0.568	1.000	0.134	-0.672	-0.824	-0.918	1.000
Serial co	Serial correlation							
0.167	0.181	0.395	0.278	0.150	-0.046	-0.052	-0.063	-0.079

Note: \1 For Peru, as quarterly data for terms of trade is only available since 1990.I, the second moments in data reported here for this variable covers only the period 1990.2-2013.IV.

Moments

Table: Second moments

Statistics	Statistics : Chile							
		Data				Mo	del	
$\Delta ln(Y)$	$\Delta ln(C)$	$\Delta ln(I)$	$\Delta TB/Y$	$\Delta ln(ToT)$	$\Delta ln(Y)$	$\Delta ln(C)$	$\Delta ln(I)$	$\Delta TB/Y$
Standard	Standard deviation (in %)							
1.287	1.333	4.471	2.084	4.974	1.285	1.656	4.789	1.812
Relative .	Relative S.D w.r.t $\Delta ln(Y)$							
1.000	1.036	3.474	1.619	3.865	1.000	1.289	3.727	1.410
Correlati	on with Δlr	$\iota(Y)$						
1.000	0.783	0.591	0.030	0.223	1.000	0.763	0.693	-0.301
Correlati	on with ΔT	B/Y						
0.030	-0.149	-0.410	1.000	0.487	-0.301	-0.721	-0.759	1.000
Serial co	Serial correlation							
0.169	0.384	0.321	0.268	0.484	-0.114	-0.087	-0.076	-0.030

Note: \1 For Peru, as quarterly data for terms of trade is only available since 1990.I, the second moments in data reported here for this variable covers only the period 1990.2-2013.IV.

Moments

Table: Second moments

Statistics	Statistics : Peru							
		Data				Mo	del	
$\Delta ln(Y)$	$\Delta ln(C)$	$\Delta ln(I)$	$\Delta TB/Y$	$\Delta ln(ToT)$	$\Delta ln(Y)$	$\Delta ln(C)$	$\Delta ln(I)$	$\Delta TB/Y$
Standard	deviation (in %)						
2.909	3.366	10.090	2.104	4.226	2.925	3.534	10.566	2.768
Relative :	Relative S.D w.r.t $\Delta ln(Y)$							
1.000	1.157	3.469	0.723	1.746	1.000	1.208	3.612	0.946
Correlati	on with Δln	$\iota(Y)$			'			
1.000	0.790	0.564	-0.276	0.139	1.000	0.888	0.871	-0.611
Correlati	on with ΔT	B/Y						
-0.276	-0.320	-0.350	1.000	0.320	-0.611	-0.878	-0.874	1.000
Serial co.	Serial correlation							
0.360	0.203	0.011	-0.051	0.303	-0.068	-0.090	-0.087	-0.084

Note: \int For Peru, as quarterly data for terms of trade is only available since 1990.I, the second moments in data reported here for this variable covers only the period 1990.2-2013.IV.

$Model \Rightarrow Data$

• The model in its loglinear form can be written, in matricial form, as:

$$\Gamma_0 \mathbf{W}_t + \Gamma_1 \mathbb{E}_t \mathbf{W}_{t+1} + \Gamma_2 \mathbf{W}_{t-1} + \Gamma_{\varepsilon} \, \varepsilon_t = 0 \tag{1}$$

where \mathbf{W}_t includes the set of predetermined and non-predetermined variables, ε_t stands for the shocks of the model, and Γ contain the parameters of the model.

• Next, after solving the modelo we obtain:

$$\mathbf{W}_t = \mathbf{A}\mathbf{W}_{t-1} + \mathbf{B}\boldsymbol{\varepsilon}_t. \tag{2}$$

• The observable variables can be writtern in terms of the state variables as follows

$$\mathbf{y}_t = \mathbf{Z}\mathbf{W}_t + \mathbf{d} + u_t \,, \tag{3}$$

where ${\bf Z}$ is conformable matrix, we are ready to represent the state-space form of our system.

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• We use quarterly data of gross domestic product (Y), consumption (C), investment (I), and the trade balance-to-GDP ratio (TBY) for Chile (1996.I to 2013.IV), Mexico (1980.I to 2013.IV) and Peru (1990.I-2013.IV). The map from observable variables to state variables in the model, given by equation in (3), is defined by:

$$\begin{array}{rcl} \Delta ln Y_t & = & y_t - y_{t-1} + x_{t-1} + lnX + \varepsilon_t^Y, \\ \Delta ln C_t & = & c_t - c_{t-1} + x_{t-1} + lnX + \varepsilon_t^C, \\ \Delta ln I_t & = & i_t - i_{t-1} + x_{t-1} + lnX + \varepsilon_t^I, \\ \Delta TBY_t & = & tby_t - tby_{t-1} + \varepsilon_t^{TBY}. \end{array}$$

• Given these state-space form of the model, the Kalman filter, jointly with a smoother, allows us to build recursively the estimates for the times series of the total factor productivity, $TFP_t = a_t A_t^{(1-\alpha)}$, based on estimates of the non observable variables a_t and X_t .

Decomposition of TFP: SVAR model

Next, we use a VAR model for TPF and terms of trade (TOT). We use this model to identify
domestic productivity shocks and terms of trade shocks, by considering that domestic
shocks do not have a long-run effects on terms of trade, as follows:

$$\left[\begin{array}{c} \Delta TOT_t \\ \Delta TFP_t \end{array}\right] = B(L) \left[\begin{array}{c} \Delta TOT_{t-1} \\ \Delta TFP_{t-1} \end{array}\right] + C_{\varepsilon} \left[\begin{array}{c} \varepsilon_t^{TOT} \\ \varepsilon_t^{TFP} \end{array}\right],$$

The identification restriction implies that domestic TFP shocks do not affect terms of trade in the long-run, therefore, the long-run impact matrix $\Theta(1)$, is restricted as follows:

$$\Theta(1) = \lim_{j \to \infty} \left(E_t - E_{t-1} \right) \begin{bmatrix} TOT_{t+j} \\ TFP_{t+j} \end{bmatrix} = \begin{bmatrix} \sum_{s=0}^{s=\infty} \theta_{11}^{(s)} & 0 \\ \sum_{s=0}^{s=\infty} \theta_{21}^{(s)} & \sum_{s=0}^{s=\infty} \theta_{22}^{(s)} . \end{bmatrix}$$

• This assumption takes into account the fact that terms of trade in this economies are mostly driven by exogenous fluctuations in commodity prices.

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TFP decomposition

The impact of terms of trade shocks are not negligible in the long run...

Table: VAR results: Variance decomposition of TFP (in %)

Horizon (k)	Cł	nile	Mexic		ico Pe	
(quarters)	$arepsilon_t^{TOT}$	$arepsilon_t^{TFP}$	$arepsilon_t^{TOT}$	$arepsilon_t^{TFP}$	ϵ_t^{TOT}	$arepsilon_t^{TFP}$
1	1.45	98.55	14.40	85.60	2.43	97.57
2	10.45	89.55	19.63	80.37	8.39	91.61
3	13.62	86.38	19.99	80.01	9.32	90.68
4	14.38	85.62	20.02	79.98	9.43	90.57
10	14.57	85.43	20.02	79.98	9.45	90.55
40	14.57	85.43	20.02	79.98	9.45	90.55

Note: Each figure at horizon k stands for the percentage of the k-quarter forecasting error explained by ε_t^{TOT} or ε_t^{TFP} shocks. For Chile the results correspond to a VAR(1), for Mexico to a VAR(2) and a VAR(1) for Peru.

Historical decomposition: Chile

... and in the short run.

Table: TFP decomposition (Average annual growth rate, in %)

	Terms of	,			
Period	Trade	DSGE estimation\1 (a)	Without ε_t^{TOT} shocks $^{\setminus 2}$ (b)	Difference (a) -(b)	
Chile, VAR(1)					
1998-2000	5.308	1.191	1.178	0.013	
2001-2007	13.086	2.848	2.408	0.440	
2008-2013	3.338	1.926	2.087	-0.161	
2008	-12.894	1.227	2.011	-0.784	
2009	0.953	-0.628	0.371	-0.999	
2010	23.308	3.092	1.885	1.208	
2011	3.922	3.165	2.974	0.191	
2012	-6.341	2.276	3.000	-0.724	
2013	-2.740	1.770	2.227	-0.457	
1998-2013	5.422	2.117	2.089	0.028	
STD(in %)	4.974	1.239	0.884		

 $^{\ ^1}$ Estimates based on stage #1. $\ ^2$ Estimates based on stage #2.

Historical decomposition: Mexico

Table: TFP decomposition (Average annual growth rate, in %)

	Terms of				
Period	Trade	DSGE estimation\1 (a)	Without ε_t^{TOT} shocks $^{\setminus 2}$ (b)	Difference (a) -(b)	
Mexico, VAR(2)					
1982-1990	-5.378	0.721	1.144	-0.423	
1991-2000	0.319	0.831	0.663	0.168	
2001-2007	2.752	1.631	1.206	0.425	
2008-2013	-0.317	0.781	0.648	0.133	
2008	1.235	0.046	-0.331	0.377	
2009	-11.063	-2.021	-0.848	-1.173	
2010	7.588	1.403	0.383	1.020	
2011	6.820	2.079	1.255	0.824	
2012	-3.640	1.534	1.612	-0.077	
2013	-0.070	0.961	0.857	0.105	
1982-2013	-1.189	0.947	0.900	0.047	
STD(in %)	4.857	1.197	0.907		

Historical decomposition: Peru

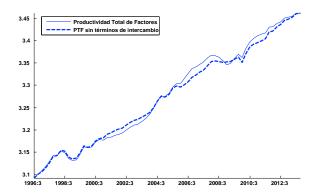
Table: TFP decomposition (Average annual growth rate, in %)

	Terms of				
Period	Trade	DSGE estimation\1 (a)	Without ε_t^{TOT} shocks $^{\setminus 2}$ (b)	Difference (a) -(b)	
Peru, VAR(1)					
1981-1990	-3.887	-2.405	n.a		
1992-2000	-2.004	2.242	2.655	-0.413	
2001-2007	8.060	3.625	2.449	1.176	
2008-2013	1.823	2.421	2.331	0.090	
2008	-14.523	3.871	5.916	-2.045	
2009	-3.095	-0.187	1.793	-1.980	
2010	18.223	5.089	2.077	3.012	
2011	5.526	3.618	2.504	1.114	
2012	-4.957	2.473	3.377	-0.904	
2013	-4.743	1.194	1.912	-0.719	
1992-2013	1.049	2.634	2.553	0.080	
STD(in %)	4.226	2.371	1.998		

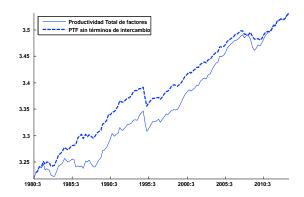
Conclusions

- Empirical results show that terms of trade shocks had indeed generated important gains in TFP for the Chilean, Mexican and Peruvian economy, in during the 2000s.
- The estimation also shows that the periods of negative terms of trade explain an slowdown in TFP growth, in particular during the recession period of 2008-2009.
- Short-run effects seem more predominant than long-run effects. In the long-run, the variance decomposition shows that the terms of trade are more important for Mexico and Chile; whereas for Peru the short-run effects seems more important.
- Extensions,
 - We can explore with the model also the impact of terms of trade shocks on investment and consumption. We require to include two unit roots in the model. We leave this for future work.

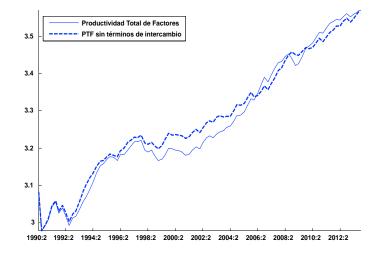
Chile: Logaritmo de la productividad total de factores (1996.Il-2013.IV)



México: Logaritmo de la productividad total de factores (1980.II-2013.IV)



Perú: Logaritmo de la productividad total de factores (1990.I-2013.IV)



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